

Stochastic Processes

MATH5835, P. Del Moral

UNSW, School of Mathematics & Statistics

Lectures Notes, No. 4

Consultations (RC 5112):

Wednesday 3.30 pm \rightsquigarrow 4.30 pm & Thursday 3.30 pm \rightsquigarrow 4.30 pm

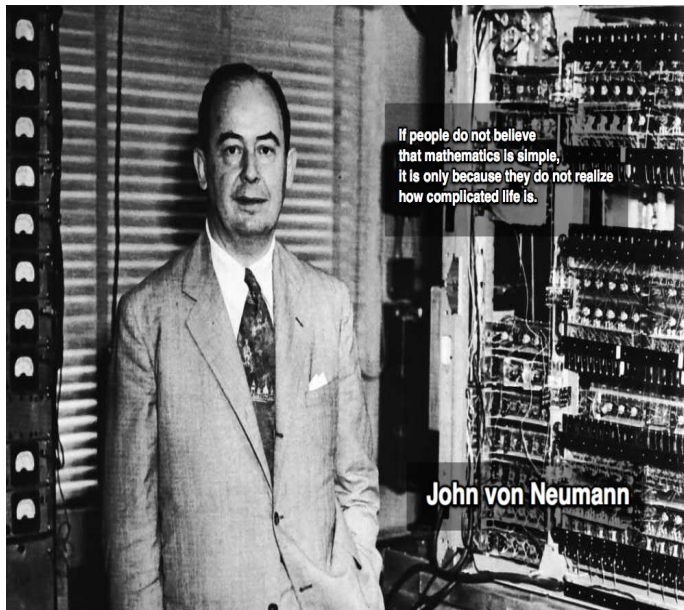
References in the slides

- ▶ **Material for research projects** \rightsquigarrow Moodle
(*Stochastic Processes and Applications* \ni variety of applications)
- ▶ **Important results**

⊂ **Assessment/Final exam** = LOGO =



Citations of the day



If people do not believe
that mathematics is simple,
It is only because they do not realize
how complicated life is.

John von Neumann

J.v.N. 1903-1957

Gauss/Legendre Regression formula (1800)



- ▶ **Linear-Gaussian estimation problem**

$$\begin{cases} X \sim N(m, \sigma^2) \\ Y = bX + V \quad (X \perp) V \sim N(0, \tau^2) \end{cases}$$

- ▶ **Solution ? $\rightsquigarrow X | Y$**

Gauss/Legendre Regression formula (1800)



▶ Linear-Gaussian estimation problem

$$\begin{cases} X \sim N(m, \sigma^2) \\ Y = bX + V \quad (X \perp) V \sim N(0, \tau^2) \end{cases}$$

▶ Solution ? $\rightsquigarrow X | Y$ is a Gaussian r.v. with

$$\mathbb{E}(X|Y) = \mathbb{E}(X) + \frac{\text{Cov}(X, Y)}{\text{Var}(Y)} (Y - \mathbb{E}(Y))$$

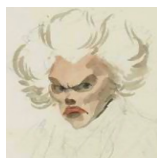
and

$$\text{Var}(X | Y) = \text{Var}(X) \left[1 - \frac{\text{Cov}(X, Y)^2}{\text{Var}(X) \text{Var}(Y)} \right]$$

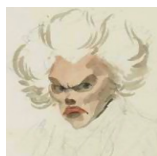
Gauss - Legendre formula (1800)



Gauss - Legendre formula (1800)



Gauss - Legendre formula (1800)



$$X \sim N(m, \sigma^2) \rightsquigarrow Y = b X + N(0, \tau^2) \Rightarrow \mathbb{E}(Y) = b m$$

↓

$$Y - \mathbb{E}(Y) = b (X - \mathbb{E}(X)) + V \Rightarrow \begin{cases} \text{Var}(Y) = b^2 \sigma^2 + \tau^2 \\ \text{Cov}(X, Y) = b \sigma^2 \end{cases}$$

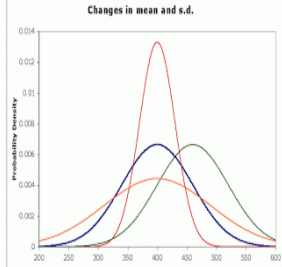
↓

$$\mathbb{E}(X|Y) = \hat{m} := \frac{\hat{\sigma}^2}{\sigma^2} m + \left(1 - \frac{\hat{\sigma}^2}{\sigma^2}\right) b^{-1} Y$$

with

$$\text{Var}(X | Y) := \hat{\sigma}^2 = \sigma^2 \left[1 - \frac{b^2 \sigma^2}{(b^2 \sigma^2 + \tau^2)}\right] = (b^2 \tau^{-2} + \sigma^{-2})^{-1}$$

Linear transformations



$$X \sim N(\hat{m}, \hat{\sigma}^2)$$

$$\Rightarrow \text{Law} \left(a X \overset{\perp}{+} N(0, \sigma^2) \right) = N(a \hat{m}, a^2 \hat{\sigma}^2 + \sigma^2)$$

Proof:

$$\begin{aligned} & (a(X - \hat{m}) + N(0, \sigma^2))^2 \\ &= a^2 \underbrace{(X - \hat{m})^2}_{\mathbb{E}(\dots) = \hat{\sigma}^2} + 2a \underbrace{(X - \hat{m}) N(0, \sigma^2)}_{\mathbb{E}(\dots) = 0} + \underbrace{(N(0, \sigma^2))^2}_{\mathbb{E}(\dots) = \sigma^2} \end{aligned}$$

The Kalman filter (1960)



Linear-Gaussian [$X_0 \sim N(m_0, \sigma_0^2) \perp W_n \sim N(0, \sigma^2) \perp V_n \sim N(0, \tau^2)$]

$$\begin{cases} X_n = a X_{n-1} + W_n & \text{Signal} \\ Y_n = b X_n + V_n & \text{Observation} \end{cases}$$

The Kalman filter (1960)



Linear-Gaussian [$X_0 \sim N(m_0, \sigma_0^2) \perp W_n \sim N(0, \sigma^2) \perp V_n \sim N(0, \tau^2)$]

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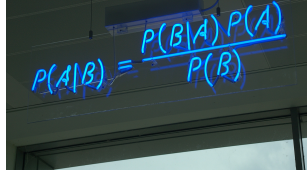
Filtering problem

$\underbrace{\text{Law}(X_n | Y_0, \dots, Y_{n-1})}_{\text{one-step predictor}}$ and $\underbrace{\text{Law}(X_n | Y_0, \dots, Y_{n-1}, Y_n)}_{\text{optimal filter}}$

\Updownarrow **Bayesian "density" notation**

" $p(x_n | y_0, \dots, y_{n-1})$ " and " $p(x_n | y_0, \dots, y_n)$ "

The Kalman filter 1/4



The first observation Y_0 :

$$X_0 \sim N(m_0, \sigma_0^2) \rightsquigarrow Y_0 = b X_0 + N(0, \tau^2)$$

First updating step = the Bayes' rule

$$p(x_0) = N(m_0, \sigma_0) \rightarrow p(x_0 | y_0) \propto p(y_0 | x_0) p(x_0) = N(\hat{m}_0, \hat{\sigma}_0)$$

with the regression formula

$$\begin{aligned}\hat{m}_0 &:= \frac{\hat{\sigma}_0^2}{\sigma_0^2} m_0 + \left(1 - \frac{\hat{\sigma}_0^2}{\sigma_0^2}\right) b^{-1} Y_0 \\ \hat{\sigma}_0^2 &= (b^2 \tau^{-2} + \sigma_0^{-2})^{-1}\end{aligned}$$

The Kalman filter 2/4

The first prediction

$$p(x_0 | y_0) \rightarrow \int_{x_0} \underbrace{p(x_1 | x_0)}_{:= \text{Law}(ax_0 + W_1)} \times \underbrace{p(x_0 | y_0)}_{:= N(\hat{m}_0, \hat{\sigma}_0^2)} dx_0 = p(x_1 | y_0)$$

⇓

$$a N(\hat{m}_0, \hat{\sigma}_0^2) + W_1 \sim N \left[\underbrace{a \hat{m}_0}_{:= m_1}, \underbrace{a^2 \hat{\sigma}_0^2 + \sigma^2}_{:= \sigma_1^2} \right] = \text{Law}(X_1 | Y_0)$$

The Kalman filter 3/4

Second updating step = the Bayes' rule

$$p(x_1 | y_0) = N(m_1, \sigma_1^2)$$

↓

$$p(x_1 | y_0, y_1) \propto p(y_1 | x_1) p(x_1 | y_0) = N(\hat{m}_1, \hat{\sigma}_1^2)$$

with the regression formula

$$\begin{aligned}\hat{m}_1 &:= \frac{\hat{\sigma}_1^2}{\sigma_1^2} m_1 + \left(1 - \frac{\hat{\sigma}_1^2}{\sigma_1^2}\right) b^{-1} Y_1 \\ \hat{\sigma}_1^2 &= (b^2 \tau^{-2} + \sigma_1^{-2})^{-1}\end{aligned}$$

The Kalman filter 4/4



The second prediction

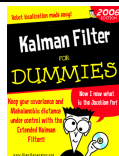
$$p(x_1|y_0, y_1) \rightarrow \int_{x_1} \underbrace{p(x_2 | x_1)}_{:=\text{Law}(ax_1+W_2)} \times \underbrace{p(x_1|y_0, y_1)}_{:=N(\hat{m}_1, \hat{\sigma}_1^2)} dx_1 = p(x_2|y_0, y_1)$$


⇓

$$a N(\hat{m}_1, \hat{\sigma}_1^2) + W_2 \sim N \left[\underbrace{a \hat{m}_1}_{:=m_2}, \underbrace{a^2 \hat{\sigma}_1^2 + \sigma^2}_{:=\sigma_2^2} \right] = \text{Law}(X_2|Y_0, Y_1)$$

AND SO ON ... = KALMAN FILTER

- ⊕ drone board (altimeter/angle observations)
- ⊕ Kalman filters in mathworks
- ⊕ Section 3.2 + Section 4.5.1 + Section 4.5.4 (& 4.5.3  )



Particle Filter  = Population dynamics model

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Initially

$$(\xi_0^1, \dots, \xi_0^N) = N \text{ i.i.d. copies of } X_0 \implies \frac{1}{N} \sum_{1 \leq i \leq N} \delta_{\xi_0^i} \simeq_{N \uparrow \infty} \text{Law}(X_0)$$

Particle Filter = Population dynamics model

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WHY??

Particle Filter = Population dynamics model

Initially

$$(\xi_0^1, \dots, \xi_0^N) = N \text{ i.i.d. copies of } X_0 \implies \frac{1}{N} \sum_{1 \leq i \leq N} \delta_{\xi_0^i} \simeq_{N \uparrow \infty} \text{Law}(X_0)$$

WHY??

The law of large numbers: $\forall f : \mathbb{R} \mapsto \mathbb{R}$ (ex.: $f = 1_{[a,b]}$)

$$\frac{1}{N} \sum_{1 \leq i \leq N} f(\xi_0^i) \simeq_{N \uparrow \infty} \mathbb{E}(f(X_0)) = \int f(x_0) p(x_0) dx_0$$

Exercise:

$$\mathbb{E} \left(\frac{1}{N} \sum_{1 \leq i \leq N} f(\xi_0^i) \right) = \dots?? \quad \mathbb{E} \left(\left[\frac{1}{N} \sum_{1 \leq i \leq N} f(\xi_0^i) - \mathbb{E}(f(X_0)) \right]^2 \right) = \dots??$$

Integral representation

Dirac measure δ_a at some point a :

$$\int f(x) \delta_a(dx) = f(a)$$

\Downarrow

$$\frac{1}{N} \sum_{1 \leq i \leq N} f(\xi_0^i) = \frac{1}{N} \sum_{1 \leq i \leq N} \int f(x_0) \delta_{\xi_0^i}(dx_0)$$

Integral representation

Dirac measure δ_a at some point a :

$$\int f(x) \delta_a(dx) = f(a)$$

\Downarrow

$$\begin{aligned} \frac{1}{N} \sum_{1 \leq i \leq N} f(\xi_0^i) &= \frac{1}{N} \sum_{1 \leq i \leq N} \int f(x_0) \delta_{\xi_0^i}(dx_0) \\ &= \int f(x_0) \left[\frac{1}{N} \sum_{1 \leq i \leq N} \delta_{\xi_0^i} \right] (dx_0) \\ &\simeq \int f(x_0) p(x_0) dx_0 \end{aligned}$$


A note on likelihood functions

$$Y_n = h_n(X_n) + N(0, \tau^2)$$

↓

$$p(y_n | x_n) = \frac{1}{\sqrt{2\pi\tau^2}} e^{-\frac{1}{2\tau^2}(y_n - h_n(x_n))^2}$$


= Large values iff $h_n(x_n)$ close to the observation y_n

First updating w.r.t. $Y_0 =$ Selection of individuals 

$$p(x_0)dx_0 \simeq \frac{1}{N} \sum_{1 \leq i \leq N} \delta_{\xi_0^i}(dx_0)$$

\Downarrow

$$p(x_0 | y_0)dx_0 = \frac{p(y_0 | x_0) p(x_0)dx_0}{\int p(y_0 | x'_0) p(x'_0)dx'_0}$$

First updating w.r.t. $Y_0 =$ Selection of individuals 

$$p(x_0)dx_0 \simeq \frac{1}{N} \sum_{1 \leq i \leq N} \delta_{\xi_0^i}(dx_0)$$

\Downarrow

$$\begin{aligned} p(x_0 | y_0)dx_0 &= \frac{p(y_0 | x_0) p(x_0)dx_0}{\int p(y_0 | x'_0) p(x'_0)dx'_0} \\ &\simeq \sum_{1 \leq i \leq N} \frac{p(y_0 | \xi_0^i)}{\sum_{1 \leq j \leq N} p(y_0 | \xi_0^j)} \delta_{\xi_0^i}(dx_0) \end{aligned}$$

\Downarrow

Sampling N r.v. $(\hat{\xi}_0^1, \dots, \hat{\xi}_0^N)$ with this law:

$$p(x_0 | y_0)dx_0 \simeq \frac{1}{N} \sum_{1 \leq i \leq N} \delta_{\hat{\xi}_0^i}(dx_0)$$

The first prediction

$$p(x_1 | y_0) := \int_{x_0} p(x_1 | x_0) \times \underbrace{p(x_0 | y_0)}_{\simeq \frac{1}{N} \sum_{1 \leq i \leq N} \delta_{\hat{\xi}_0^i}} dx_0$$

\simeq Law of the first state of a Markov chain
starting at some $\hat{\xi}_0^i$ with probability $1/N$, with $i = 1, \dots, N$



Sample N transitions $X_0 \rightsquigarrow X_1$ starting at $\hat{\xi}_0^1, \dots, \hat{\xi}_0^N$

$$p(x_1 | y_0) dx_1 \simeq \frac{1}{N} \sum_{1 \leq i \leq N} \delta_{\hat{\xi}_0^i}(dx_1) \quad \text{with} \quad \xi_1^i \sim p(x_1 | \hat{\xi}_0^i)$$

Second updating w.r.t. $Y_1 =$ Selection of individuals 

$$p(x_1 | y_0) dx_1 \simeq \frac{1}{N} \sum_{1 \leq i \leq N} \delta_{\xi_i}(dx_1)$$

\Downarrow

$$p(x_1 | y_0, y_1) dx_1 = \frac{p(y_1 | x_1) p(x_1 | y_0) dx_1}{\int p(y_1 | x'_1) p(x'_1 | y_0) dx'_1}$$

Second updating w.r.t. $Y_1 =$ Selection of individuals 

$$p(x_1 | y_0) dx_1 \simeq \frac{1}{N} \sum_{1 \leq i \leq N} \delta_{\xi_1^i}(dx_1)$$

⇓

$$\begin{aligned} p(x_1 | y_0, y_1) dx_1 &= \frac{p(y_1 | x_1) p(x_1 | y_0) dx_1}{\int p(y_1 | x'_1) p(x'_1 | y_0) dx'_1} \\ &\simeq \sum_{1 \leq i \leq N} \frac{p(y_1 | \xi_1^i)}{\sum_{1 \leq j \leq N} p(y_1 | \xi_1^j)} \delta_{\xi_1^i}(dx_1) \end{aligned}$$

⇓

Sampling N r.v. $(\hat{\xi}_1^1, \dots, \hat{\xi}_1^N)$ with this law:

$$p(x_1 | y_0, y_1) dx_1 \simeq \frac{1}{N} \sum_{1 \leq i \leq N} \delta_{\hat{\xi}_1^i}(dx_1)$$

The second prediction

$$p(x_2|y_0, y_1) := \int_{x_1} p(x_2 | x_1) \times \underbrace{p(x_1|y_0, y_1)}_{\simeq \frac{1}{N} \sum_{1 \leq i \leq N} \delta_{\hat{\xi}_1^i}} dx_1$$

\simeq Law of the second state X_2 of a Markov chain
given $X_1 = \hat{\xi}_1^i$ with probability $1/N$.



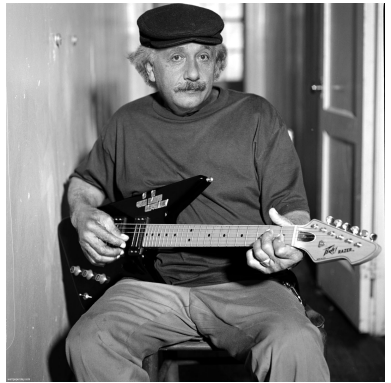
Sample N transitions $X_1 \rightsquigarrow X_2$ starting at $\hat{\xi}_1^1, \dots, \hat{\xi}_1^N$

$$p(x_2|y_0, y_1) dx_2 \simeq \frac{1}{N} \sum_{1 \leq i \leq N} \delta_{\xi_2^i}(dx_2) \quad \text{with} \quad \xi_2^i \sim p(x_2 | \hat{\xi}_1^i)$$

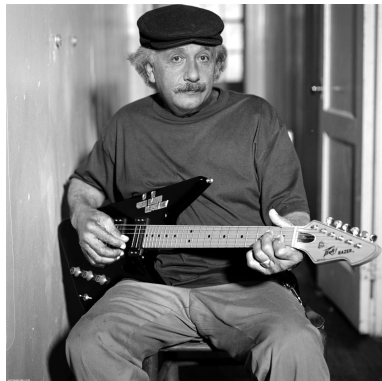
AND SO ON ... = PARTICLE FILTER (section 3.2)

The lost equation - Introduction to Brownian motion

Second citation of the day

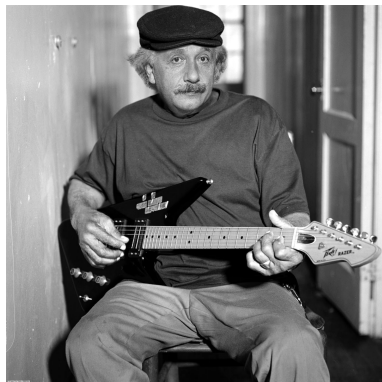


Second citation of the day



Since the mathematicians have invaded the theory of relativity, I do not understand it myself anymore. – *Albert Einstein (1879-1955)*

Second citation of the day



Since the mathematicians have invaded the theory of relativity, I do not understand it myself anymore. – *Albert Einstein (1879-1955)*

His seventh-grade teacher, exasperated by young Albert's insubordination, famously told Einstein he "would never get anywhere in life".
[A. Calaprice, *The New Quotable Einstein* (Princeton U.P. 2005)].

Brownian motion B_t or W_t

Discrete time version : "dt" time steps \oplus fair coin tossing

$$W_t := W_{t-dt} + \begin{cases} +\sqrt{dt} & \text{if Heads} \\ -\sqrt{dt} & \text{if Tails} \end{cases}$$

or
$$W_t := W_{t-dt} + \sqrt{dt} \times N(0, 1)$$

\oplus 2d BM (video) [BM-scale.avi/mp4](#)



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$dt = 10^{-10000000}?? \simeq_{dt \sim 0}$ **Continuous time \oplus stochastic calculus**